

Derivation of the fact that the distribution of the sample mean is the normal distribution

Consider taking n samples from a population characterized by mean, μ , and variance, σ^2 . The sample mean is given by \bar{x} .

We define a moment generating function for a continuous PDF to be:

$$M_x(t) = \mu_{e^{tx}} = E[e^{tx}] = \int_{-\infty}^{\infty} e^{tx} f(x) dx$$

From this definition, and using the rules of linear operation, we can show that

$$M_{x+a}(t) = \mu_{e^{t(x+a)}} = E[e^{t(x+a)}] = \int_{-\infty}^{\infty} e^{t(x+a)} f(x) dx = \int_{-\infty}^{\infty} e^{tx} e^{ta} f(x) dx = e^{ta} \int_{-\infty}^{\infty} e^{tx} f(x) dx$$

$$M_{x+a}(t) = e^{ta} M_x(t) \quad (1)$$

$$M_{ax}(t) = \mu_{e^{tax}} = E[e^{tax}] = \int_{-\infty}^{\infty} e^{tax} f(x) dx = M_x(at)$$

$$M_{ax}(t) = M_x(at) \quad (2)$$

$$M_{x_1+x_2+x_3+\dots+x_n}(t) = \mu_{e^{t(x_1+x_2+x_3+\dots+x_n)}} = E[e^{t(x_1+x_2+x_3+\dots+x_n)}] = \int_{-\infty}^{\infty} e^{t(x_1+x_2+x_3+\dots+x_n)} f(x) dx$$

$$M_{x_1+x_2+x_3+\dots+x_n}(t) = M_{x_1}(t) M_{x_2}(t) M_{x_3}(t) \dots M_{x_n}(t) \quad (3)$$

So that

$$M_{(\bar{x}-\mu)/(\sigma/\sqrt{n})}(t) = \mu_{e^{t[(\bar{x}-\mu)/(\sigma/\sqrt{n})]}} = E[e^{t[(\bar{x}-\mu)/(\sigma/\sqrt{n})]}] = \int_{-\infty}^{\infty} e^{t[(\bar{x}-\mu)/(\sigma/\sqrt{n})]} f(\bar{x}) d\bar{x}$$

$$M_{(\bar{x}-\mu)/(\sigma/\sqrt{n})}(t) = \int_{-\infty}^{\infty} e^{t\bar{x}/(\sigma/\sqrt{n})} e^{-t\mu/(\sigma/\sqrt{n})} f(\bar{x}) d\bar{x} = e^{-t\mu\sqrt{n}/\sigma} \int_{-\infty}^{\infty} e^{t\bar{x}\sqrt{n}/\sigma} f(\bar{x}) d\bar{x}$$

$$M_{(\bar{x}-\mu)/(\sigma/\sqrt{n})}(t) = e^{-t\mu\sqrt{n}/\sigma} M_{\bar{x}}\left(\frac{t\sqrt{n}}{\sigma}\right)$$

Now consider that $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$, so

$$M_{\bar{x}}\left(\frac{t\sqrt{n}}{\sigma}\right) = M_{\sum_{i=1}^n X_i}\left(\frac{t\sqrt{n}}{\sigma}\right) = \int_{-\infty}^{\infty} e^{t\sqrt{n}/\sigma \sum_{i=1}^n X_i} f(x) dx = \int_{-\infty}^{\infty} e^{t/(\sigma\sqrt{n}) \sum_{i=1}^n X_i} f(x) dx$$

$$M_{\bar{x}}(t\sqrt{n}/\sigma) = \int_{-\infty}^{\infty} e^{t/(\sigma\sqrt{n}) \sum_{i=1}^n X_i} f(x) dx = M_{X_1}\left(\frac{t}{\sigma\sqrt{n}}\right) M_{X_2}\left(\frac{t}{\sigma\sqrt{n}}\right) \dots M_{X_n}\left(\frac{t}{\sigma\sqrt{n}}\right)$$

Since there isn't anything intrinsic that distinguishes one X_i from another, we can write

$$M_{\bar{x}}(t\sqrt{n}/\sigma) = \left[M_x\left(\frac{t}{\sigma\sqrt{n}}\right) \right]^n$$

If we substitute this back into our original equation

$$M_{\frac{(\bar{x}-\mu)}{(\sigma/\sqrt{n})}}(t) = e^{-t\mu\sqrt{n}/\sigma} M_{\bar{x}}\left(\frac{t\sqrt{n}}{\sigma}\right) = e^{-t\mu\sqrt{n}/\sigma} \left[M_x\left(\frac{t}{\sigma\sqrt{n}}\right) \right]^n$$

Take the natural log of both sides:

$$\ln \left[M_{\frac{(\bar{x}-\mu)}{(\sigma/\sqrt{n})}}(t) \right] = \frac{-t\mu\sqrt{n}}{\sigma} + n \ln \left[M_x\left(\frac{t}{\sigma\sqrt{n}}\right) \right]$$

Expand $M_x\left(\frac{t}{\sigma\sqrt{n}}\right)$ as an infinite series in powers of t about $t=0$.

$$M_x\left(\frac{t}{\sigma\sqrt{n}}\right) = 1 + v_1 t + v_2 \frac{t^2}{2!} + v_3 \frac{t^3}{3!} + \dots + v_r \frac{t^r}{r!} + \dots$$

where

$$v_i = \left. \frac{d^i M_x\left(\frac{t}{\sigma\sqrt{n}}\right)}{dt^i} \right|_{t=0}$$

We can write this as

$$M_x\left(\frac{t}{\sigma\sqrt{n}}\right) = 1 + v(t)$$

where $v(t)$ is an infinite series in t . For very large sample sizes, n

$$\lim_{n \rightarrow \infty} \ln \left[M_{\frac{(\bar{x}-\mu)}{(\sigma/\sqrt{n})}}(t) \right] = \lim_{n \rightarrow \infty} \ln[1 + v(t)] = \frac{t^2}{2}$$

This can be shown by expanding the natural log in a Mclaurin series. For the present purposes, we will take this step given above on faith. Then, we have

$$\lim_{n \rightarrow \infty} M_{\frac{(\bar{x}-\mu)}{(\sigma/\sqrt{n})}}(t) = e^{\frac{t^2}{2}}$$

So the first moment of $Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}}$ in the limit of large n is $e^{\frac{t^2}{2}}$

Well, let's find what the moment of the random variable, z , would be if it follows the normal distribution. The PDF of the normal distribution is

$$f(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

$$M_x(t) = \mu_{e^{tx}} = E[e^{tx}] = \int_{-\infty}^{\infty} e^{tx} f(x) dx = \int_{-\infty}^{\infty} e^{tx} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} dx$$

$$M_x(t) = \int_{-\infty}^{\infty} e^{tx} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{-x^2 + 2x\mu - \mu^2}{2\sigma^2}} dx = \int_{-\infty}^{\infty} e^{\frac{2tx\sigma^2}{2\sigma^2}} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{-x^2 + 2x\mu - \mu^2}{2\sigma^2}} dx$$

$$M_x(t) = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{\frac{2tx\sigma^2 - x^2 + 2x\mu - \mu^2}{2\sigma^2}} dx = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{\frac{-x^2 + 2(t\sigma^2 + \mu)x - \mu^2}{2\sigma^2}} dx$$

Complete the square in the exponent:

$$-x^2 + 2(t\sigma^2 + \mu)x - \mu^2 = [x - (t\sigma^2 + \mu)]^2 - 2\mu t\sigma^2 - t^2\sigma^4$$

$$M_x(t) = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{\frac{[x - (t\sigma^2 + \mu)]^2 - 2\mu t\sigma^2 - t^2\sigma^4}{2\sigma^2}} dx = e^{\mu t + \frac{t^2\sigma^2}{2}} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{[x - (t\sigma^2 + \mu)]^2}{2\sigma^2}} dx$$

$$\text{Let } w = \frac{[x - (t\sigma^2 + \mu)]}{\sigma} \text{ so that } dw = \frac{dx}{\sigma} \text{ and}$$

$$M_x(t) = e^{\mu t + \frac{t^2 \sigma^2}{2}} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{w^2}{2}} dw = e^{\mu t + \frac{t^2 \sigma^2}{2}} (1) = e^{\mu t + \frac{t^2 \sigma^2}{2}}$$

So that the first moment generating function of the standard normal PDF is

$$M_x(t) = e^{\frac{t^2}{2}}$$

If we compare this moment generating function with that obtained for

$$\lim_{n \rightarrow \infty} M_{\frac{(\bar{x} - \mu)}{(\sigma/\sqrt{n})}}(t) = e^{\frac{t^2}{2}}$$

we find that they are the same in the limit of large n . Since there is a one-to-one correspondence between PDFs and moment-generating functions, we see that the PDF for

$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}}$ is the standard normal PDF.